

Lecture #23

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So, if $\hat{f}_h(x) = \frac{1}{n} \cdot \sum_{i=1}^n \frac{1}{h} K\left(\frac{x-x_i}{h}\right)$, can we describe what happens as the bandwidth h varies?

Well, recall that K has integral $1 = \int_{\mathbb{R}} K(x) dx = 1$. This suggests

that K concentrates near the origin. So, let's suppose that in fact K is compactly supported, i.e. $\exists \delta > 0$ s.t. $\int_{-\delta}^{\delta} K(x) dx = 1$. In the event K is non-compactly supported, $\forall \epsilon > 0 \exists \delta > 0$ s.t. $\int_{-\delta}^{\delta} K(x) dx \geq 1 - \epsilon$, so

The same is true Spiritually, if not literally.

So, if $\int_{-\delta}^{\delta} K(x) dx = 1$, then $\int_{-x_i}^{x_i} K(x-x_i) dx = 1$

$$\Rightarrow \int_{-[x_i] \cdot h}^{[x_i] \cdot h} \frac{1}{h} K\left(\frac{x-x_i}{h}\right) dx = 1$$

We conclude that $\frac{1}{h} K\left(\frac{x-x_i}{h}\right)$ concentrates around x_i for h small

($\frac{1}{h} K\left(\frac{x-x_i}{h}\right)$ is tall and skinny), while $\frac{1}{h} K\left(\frac{x-x_i}{h}\right)$ is diffuse for h large ($\frac{1}{h} K\left(\frac{x-x_i}{h}\right)$ is short and fat). You will see this in the

Supplemental Question 1 on HW 11.

As in the case of histogram estimation, the tunable parameter h can be set to optimally balance between bias and variance: (2)

Theorem (Bias-Variance Tradeoff for KDE): Let K be a kernel and let $\hat{f}_n(x) = \frac{1}{n} \sum_{i=1}^n \frac{1}{h} K\left(\frac{x-x_i}{h}\right)$ be a kernel density estimate for an underlying density $f(x)$, based on sample $\{x_i\}_{i=1}^n$ and bandwidth $h > 0$.

(i) The pointwise bias of $\hat{f}_n(x)$ is $B(x) = \mathbb{E}(\hat{f}_n(x)) - f(x) \approx \frac{1}{2} \sigma_K^2 h^2 \cdot f''(x)$,

where $\sigma_K^2 = \int_{\mathbb{R}} \frac{1}{y} K(y) dy$.

(ii) The pointwise variance of $\hat{f}_n(x)$ is $\text{Var}(x) = \mathbb{E}\left[\left(\mathbb{E}(\hat{f}_n(x)) - \hat{f}_n(x)\right)^2\right] \approx \frac{1}{nh} \left[\int_{\mathbb{R}} K^2(y) dy \right] \cdot f(x)$

(iii) The overall MSE for $\hat{f}_n(x)$ is $\text{MSE} = \int_{\mathbb{R}} [B(x)^2 + \text{Var}(x)] dx \approx \frac{1}{4} \sigma_K^4 h^4 \int_{\mathbb{R}} [f''(x)]^2 dx + \frac{1}{nh} \int_{\mathbb{R}} K^2(y) dy$

Proof: To see (i), notice

$$E(\hat{f}_n(x))$$

$$= E\left(\frac{1}{n} \sum_{i=1}^n \frac{1}{h} f\left(\frac{x-x_i}{h}\right)\right)$$

$$= \frac{1}{n} \sum_{i=1}^n E\left(\frac{1}{h} K\left(\frac{x-x_i}{h}\right)\right)$$

$$= \frac{1}{n} \sum_{i=1}^n \int_{\mathbb{R}} \frac{1}{h} K\left(\frac{x-x_i}{h}\right) f(x_i) dx_i$$

C.O.V. : $y = \frac{x-x_i}{h}$
 $dy = -\frac{1}{h} dx_i$
 $dx_i = -h \cdot dy$
 $x_i = x - hy$

$$= \frac{1}{n} \sum_{i=1}^n \int_{\mathbb{R}} K(y) f(x - hy) dy$$

$$= \frac{1}{n} \sum_{i=1}^n \int_{\mathbb{R}} K(y) \cdot [f(x) - hy f'(x) + \frac{1}{2} [hy]^2 f''(x) - \dots]$$

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$$= \frac{1}{n} \sum_{i=1}^n \left[f(x) \int_{\mathbb{R}} K(y) dy - h f'(x) \int_{\mathbb{R}} y K(y) dy \right]$$

$$+ \frac{1}{2} h^2 f''(x) \int_{\mathbb{R}} y^2 K(y) dy - \text{Higher-Order Terms}$$

$$= \frac{1}{n} \sum_{i=1}^n \left[f(x) + \frac{1}{2} h^2 f''(x) \int_{\mathbb{R}} y^2 K(y) dy - \dots \right]$$

ignore H.O.T. in h

$$\approx f(x) + \frac{1}{2} h^2 f''(x) \int_{\mathbb{R}} y^2 K(y) dy, \text{ as desired.}$$

To see (ii), we proceed similarly: $\text{Var}(X)$ (9)

$$= \text{Var}\left(\frac{1}{n} \sum_{i=1}^n \frac{1}{h} \mathcal{K}\left(\frac{x-x_i}{h}\right)\right)$$

$$= \frac{1}{n^2} \cdot \frac{1}{h^2} \text{Var}\left(\sum_{i=1}^n \mathcal{K}\left(\frac{x-x_i}{h}\right)\right)$$

$$= \frac{1}{n} \cdot \frac{1}{h^2} \cdot \text{Var}\left(\mathcal{K}\left(\frac{x-x_i}{h}\right)\right)$$

$$= \frac{1}{n} \cdot \frac{1}{h^2} \cdot \mathbb{E}\left(\mathcal{K}\left(\frac{x-x_i}{h}\right)^2\right) - \frac{1}{n} \cdot \frac{1}{h^2} \cdot \mathbb{E}\left(\mathcal{K}\left(\frac{x-x_i}{h}\right)\right)^2$$

$$= \frac{1}{n} \cdot \frac{1}{h^2} \int_{\mathbb{R}} \mathcal{K}^2\left(\frac{x-x_i}{h}\right) f(x_i) dx_i - \frac{1}{n} \cdot \frac{1}{h^2} \left[\int_{\mathbb{R}} \mathcal{K}\left(\frac{x-x_i}{h}\right) f(x_i) dx_i \right]^2$$

$y = \frac{x-x_i}{h}$
as in (i)

$$= \frac{1}{n} \cdot \frac{1}{h^2} \int_{\mathbb{R}} h \cdot \mathcal{K}^2(y) \cdot f(x-hy) dy - \frac{1}{n} \cdot \frac{1}{h^2} \left[\int_{\mathbb{R}} h \cdot \mathcal{K}(y) \cdot f(x-hy) dy \right]^2$$

$$= \frac{1}{n} \cdot \frac{1}{h} \int_{\mathbb{R}} \mathcal{K}^2(y) \cdot f(x-hy) dy - \frac{1}{n} \left[\int_{\mathbb{R}} \mathcal{K}(y) f(x-hy) dy \right]^2 \quad (\text{A})$$

Taylor expanding $f(x-hy) \stackrel{!}{=} f(x) - h[\text{stuff}]$ gives

$$\stackrel{!}{=} \frac{1}{nh} \int_{\mathbb{R}} \mathcal{K}^2(y) [f(x) + \text{H.O.T.}] dy - \frac{1}{n} \int_{\mathbb{R}} \mathcal{K}(y) [f(x) - \text{H.O.T.}] dy$$

$$= \frac{1}{nh} \cdot f(x) \cdot \int_{\mathbb{R}} \mathcal{K}^2(y) dy + \text{H.O.T. in } h$$

$$\approx \frac{1}{nh} f(x) \cdot \int_{\mathbb{R}} \mathcal{K}^2(y) dy$$

(iii) is immediate by integrating (i), (ii).

So, if $MSE \approx \frac{1}{4} \sigma_K^4 h^4 \int_{\mathbb{R}} [f''(x)]^2 dx + \frac{1}{nh} \int_{\mathbb{R}} K(x) dx,$ (5)
then the optimal h is $h^* = C \cdot n^{-\frac{1}{5}}$ for a constant C depending on f, K
but independent of n .

With such an h^* , $MSE \sim n^{-\frac{4}{5}}$, which goes to 0 as $n \rightarrow \infty$ faster
than the analogous estimate for histograms, which was $n^{-\frac{2}{3}}$.